

Probability Rules

- Seen how to work with samples from prob functions
- Seen how to summarize them.
- Need to formally manipulate probability distros

When we discuss a generic probability distro,

write

$$P(X)$$

samples

$$x_i \sim P(X)$$

Some distros

Sometimes

$$\text{prob}(X)$$

If X is discrete, for example True/False; Integers,
 $P(X)$ has no units, $\& 0 \leq P(X) \leq 1$.

~~If continuous~~ $[P(X)]$ ~~is~~

"Or" $P(X) + P(\neg X) = 1$

normalized over all possibilities.

Continuous:

$$[P(X)] = \frac{1}{[X]}$$

Normalization: $\int_{\mathbb{X}} P(X) dX = 1$

Note: $P(X)$ can be >1 in narrow regions

- $P(X)$ is a probability density or pdf

High prob density

$\not\rightarrow$ not high probability

Joint Probability

Can talk about more than 1 var:

$$P(X, Y) \text{ or } P(\vec{X}) = P(x_1, x_2, x_3, x_4)$$

discrete: X "AND" Y at same time

continuous: multi-dimensional PDF

Conditional Probability

$$P(X | Y)$$

given

probability X is true given Y is true
or
probability X given Y is fixed to some value

Product Rule

$$P(X, Y) = P(X|Y)P(Y)$$

works the other way too

$$P(X, Y) = P(Y|X)P(X)$$

↖ probability of X regardless of Y

Also works for conditional probs:

$$P(X, Y | Z) = P(X|Y, Z)P(Y|Z)$$

Bayes Rule - Put product rule together

$$P(X|Y) = \frac{P(Y|X)P(X)}{P(Y)}$$

Independence

$$P(X, Y) = P(X|Y)P(Y) = P(X)P(Y)$$

$\downarrow P(X)$

Marginalization - When we don't know a variable's value, or don't care

$$P(X) = \int_{\mathcal{Y}} P(X, Y) dY = \int_{\mathcal{Y}} P(X|Y) P(Y) dY$$

discrete:

$$P(X) = \sum_Y P(X|Y) P(Y)$$

[Supernovae Example] Tail: $P(>X) = \int_X^\infty P(X) dX$

[Prisoner Example]

Changing Variables

$$P(X) dx = P(Y) dy$$

$$P(X) = P(Y) \left| \frac{dy}{dx} \right|$$

Supernova

$$P(I_a) = 0.3 \quad \text{probability of type Ia}$$

$$P(I_b) = 0.2 \quad \text{probability of type Ib}$$

$$P(II) = 0.5$$

Note: - Discrete variable w/ 3 options

- Normalization satisfied

$$P(R) = 0.3 \quad \text{Probability of radio emission detected from SN, regardless of type}$$

- 2 option variable

$$P(II|R) = 0.75 \quad \text{- Probability that SN is type II GIVEN Radio detected}$$

$$P(I_b|R) = 0.25 \quad \text{- Probability that SN is type Ib}$$

What is $P(I_a|R) = ?$ 0. → Normalization rule!

3 Questions, with neighbors

1.) $P(R | I_b)$ probability that radio is detected
GIVEN that it is I_b

2.) $P(R | I_a \text{ or } I_b)$ probability of Radio Given type I

3.) $P(I_a | -R)$ probability of I_a given no radio

Answers:

$$1.) P(R | I_b) = \frac{P(I_b | R) P(R)}{P(I_b)} = \frac{(0.25)(0.3)}{0.2} = 0.375$$

$$\begin{aligned} 2.) P(R | I_a \text{ or } I_b) &= \frac{P(I_a \text{ or } I_b | R) P(R)}{P(I_a \text{ or } I_b)} \\ &= \frac{P(I_a | R) P(R)}{P(I_a \text{ or } I_b)} + \frac{P(I_b | R) P(R)}{P(I_a \text{ or } I_b)} = \frac{(0.25)(0.3)}{0.5} \\ &= 0.15 \end{aligned}$$

Note: $P(R | I_a \text{ or } I_b) \neq P(R | I_a) + P(R | I_b)$

~~3.)~~

$$\begin{aligned} 3.) P(I_a | -R) &= \frac{P(-R | I_a) P(I_a)}{1 - P(R)} , P(-R | I_a) = 1 - P(R | I_a) \\ &= \frac{1(0.3)}{0.7} = 0.43 \end{aligned}$$

[Notebook Demo of Answers]

3 Prisoners Dilem Problem [For Piazza]

3 prisoners in jail: A, B, C

One has randomly been selected for release.

The warden knows, but cannot say.

A asks the warden, "Tell me if B or C will be released"

"If B is to be released, tell me not C;

If C, then not B;

if me, then flip a coin and tell me which one is not released"

Warden says "Not B; after some thought for a coin flip"

A thinks his chances are now 50/50 for release, between A, C.

B A tells C, C smirks thinks he now has $\frac{2}{3}$ chance.

Who is right?

A, B, C = released

$$P(A) = P(B) = P(C) = \frac{1}{3}$$

b = told not B, c = told not C

$$P(A|b) = \frac{P(b|A)P(A)}{P(b)} =$$

$$P(b|A) = \frac{1}{2}, \quad P(b|B) = 0, \quad P(b|C) = 1$$

$$\Rightarrow P(A|b) = \frac{P(b|A)}{P(b)}$$

$$P(b) = P(b|A)P(A) + P(b|B)P(B) + P(b|C)P(C)$$

$$\Rightarrow P(A|b) = \frac{\frac{1}{2} \cdot \frac{1}{3}}{\frac{1}{2} \cdot \frac{1}{3} + 0 + \frac{1}{3}} = \frac{\frac{1}{6}}{\frac{3}{6}} = \frac{1}{3} !$$

A gets no information about his fate, but C does!

Maximum Likelihood

or "Maximum a Posteriori Solution"

$$P(\vec{x}) \rightarrow \frac{dP}{d\vec{x}} = 0 \text{ is a local max or min.}$$

Imagine we have a model $f(\vec{x}; \vec{\alpha})$
 independent parameters.
 variable

We have measured
 data \vec{y} at \vec{x} .

We want to know

$$P(\vec{\alpha} | \vec{y}) = P(\vec{y} | \vec{\alpha}) \frac{P(\vec{\alpha})}{P(\vec{y})}$$

probability of parameters after seeing data

Bayes Rule

$P(\vec{y} | \vec{\alpha})$ usually far easier to write down
 "Likelihood"

$P(\vec{\alpha})$ prior - what we think params are
 w/o seeing any data
 or \rightarrow seeing all possible datasets

$P(\vec{y})$ evidence - normalization factor
 \hookrightarrow important, hard to calculate.

\hookrightarrow come back to this.

Let's say all the data is independent.

$$P(\vec{\alpha} | \vec{y}) = \frac{P(\vec{\alpha}) \prod_i P(y_i | \vec{\alpha})}{P(\vec{y})}.$$

Let's also say $P(y_i | \vec{\alpha})$ is Gaussian, and we know the noise σ_i

$$P(y_i | \vec{\alpha}) = \frac{1}{\sqrt{2\pi\sigma_i^2}} e^{-\frac{1}{2} \left(\frac{(y_i - f(x_i; \vec{\alpha}))^2}{\sigma_i^2} \right)}$$

What is $P(\vec{\alpha})$? Might be bad, but $P(\vec{\alpha}) = U(\vec{\alpha}) = \frac{1}{N}$
 constant

$P(\vec{\alpha} | \vec{y})$ getting messy, hard to calc. We can always take its Log:

$$\begin{aligned}\ln P(\vec{\alpha} | \vec{y}) &= \sum_i \ln(P(y_i | \vec{\alpha})) + K \\ &\approx \sum_i \left[-\frac{1}{2} \left(\frac{y_i - f(x_i; \vec{\alpha})}{\sigma_i} \right)^2 - \ln(\sqrt{2\pi} \sigma_i) \right] + K\end{aligned}$$

That looks familiar!

$$\begin{aligned}X^2 &\equiv \sum_i \left(\frac{y_i - f(x_i; \vec{\alpha})}{\sigma_i} \right)^2 \\ \Rightarrow \ln P(\vec{\alpha} | \vec{y}) &\propto -\frac{1}{2} X^2 + K \quad \text{wrt } \vec{\alpha}.\end{aligned}$$

So maximizing $\ln P$ what we do to find the mode
of the distro
is same as minimizing X^2 .

But what assumptions?

$\left\{ \begin{array}{l} \text{i.i.d. data} \\ \text{Gaussian Noise} \\ P(\vec{\alpha}) = \text{Uniform.} \end{array} \right.$
Often not the case!

What if we had 1D Gaussian data? Want μ ?

$$\Rightarrow \ln P(\mu | \vec{x}) \propto -\frac{1}{2} \sum_i \left(\frac{x_i - \mu}{\sigma} \right)^2 \quad \checkmark \text{num datapoints}$$

$$\frac{d \ln P}{d \mu} = + \sum_i x_i - N\mu = \sum_i x_i - N\mu = 0$$

$$\Rightarrow \mu = \frac{1}{N} \sum_i x_i \quad \text{Max like estimate of } \mu, \text{ } \checkmark \text{ is sample mean!}$$